

controlled trial testing dose–response effects of aerobic exercise (22). In addition, we assessed if the doses of exercise would alter the magnitude of change in $\dot{V}O_{2\max}$ enough to reflect predicted risk reduction in cancer mortality in our study population of women at high risk for breast cancer.

METHODS

Study recruitment and intervention. The WISER Sister study was a randomized controlled three group parallel-arm trial designed to compare the effect of two doses of aerobic exercise training to a control group in healthy premenopausal women at elevated risk of breast cancer over 5 menstrual cycles (22). Women recruited into the WISER Sister study 1) were ≥ 18 yr, 2) were eumenorrheic, 3) were nonsmokers, 4) had a body mass index (BMI) ranging from 18 to $50 \text{ kg}\cdot\text{m}^{-2}$, 5) had a $>18\%$ lifetime risk of developing breast cancer according to the Gail or Claus prediction models (23,24), 6) had a known deleterious mutation (e.g., *BRCA1* or *BRCA2*) or documentation of a family member with the mutation, and 7) were not currently doing more than 75 min of aerobic exercise per week (22). Further details of recruitment and screening are provided in previous papers (22–25).

One hundred and thirty-nine women were randomly assigned to one of the following groups: control (CONT, $n = 46$), low-dose exercise (LOW, $n = 45$), or high-dose exercise (HIGH, $n = 48$). To account for potential confounders, participants were stratified by baseline BMI (BMI < 30 vs $\geq 30 \text{ kg}\cdot\text{m}^{-2}$) and gynecological age at start of menses (< 10 vs ≥ 10). In-home treadmills (model 5.65; Smooth Fitness, King of Prussia, PA) were provided to study participants to complete the exercise training sessions. Before randomization, all participants were provided an in-person orientation on the exercise intervention from a certified personal trainer. The trainer also provided ongoing support and monitoring of participants' adherence to the protocol. Participants randomized to treatment groups were asked to wear a downloadable Polar Heart Rate monitor (U.S. model RS400; Polar Electro Inc., Lake Success, NY) and fill out an exercise log for the purpose of objectively monitoring their exercise adherence. Heart rate data were downloaded by research staff and analyzed to document objective monitoring of exercise adherence.

Exercise intensity was set at 65%–70% of age-predicted maximum heart rate (MHR) for the first 4 wk of exercise (where age-predicted MHR = $220 - \text{age}$), which progressed to 70%–80% of MHR for the remainder of the study. Exercise intensity did not differ between the low- and the high-dose groups. Only those minutes of exercise completed within the target heart rate zone counted toward total minutes for the week. Exercise adherence was defined as the proportion of total prescribed minutes completed for the full study.

Women in the low-dose group completed $150 \text{ min}\cdot\text{wk}^{-1}$ of aerobic exercise, whereas women in the high-dose group were asked to increase exercise duration by 20–25 min every 2 wk until they reached 300 min of aerobic exercise per week. Participants randomized into the CONT group were asked to

maintain their usual level of physical activity and to not engage in any new exercise program during study participation. Strength and flexibility training were not included within this study. During the study period, 14 participants were lost to follow-up (did not return for follow-up exercise test), leaving 125 participants available for analysis (CONT, $n = 47$; LOW, $n = 39$; HIGH, $n = 39$). The study was approved by the University of Pennsylvania Institutional Review Board, and no study activities occurred before written informed participant consent.

Categorization of responders and nonresponders.

Participants' aerobic fitness level was assessed at baseline and follow-up using a maximal treadmill test, the Bruce protocol. Exercise testing was conducted 6–10 d after start of menstrual cycle. Exercise testing technicians were blinded to the intervention group allocation. $\dot{V}O_{2\max}$ was estimated using the method by Pollock et al. for maximal exercise testing with the Bruce protocol ($\dot{V}O_{2\max} = 0.073$ (time in seconds on the Bruce protocol) $- 3.9$) (26).

To categorize responders and nonresponders, we 1) determined twice the typical error ($2 \times \text{TE}$) of maximal exercise testing for $\dot{V}O_{2\max}$ estimation (27–31) and 2) defined clinic effectiveness (CE) for change in $\dot{V}O_{2\max}$ (10). TE is an adaptable form of within-subject variation; the TE of measurement reflects the reliability of measurement within groups in repeated trials, or the coefficient of variation (32). The TE for $\dot{V}O_{2\max}$ was calculated using the change from PRE to POST in the CONT group using the following equation (32):

$$\text{TE} = \frac{\text{SD}_{\text{diff}}}{\sqrt{2}}$$

where SD_{diff} is the SD of the $\dot{V}O_{2\max}$ difference in the CONT group, which estimates measurement error and the within-subject variability caused by changes in behavioral or environmental factors during the intervention period (18). This value was multiplied by two and expressed as a percentage of the measurement's mean (33). We calculated a $2 \times \text{TE}$ of $4.09 \text{ mL}\cdot\text{min}^{-1}\cdot\text{kg}^{-1}$.

Clinic effectiveness (CE) was used as another threshold to categorize exercise “responders” and “nonresponders.” CE is the minimally important difference between two groups and represents the improvement considered worthwhile by a given treatment (34). In our study, the CE for $\dot{V}O_{2\max}$ was set at $1 \text{ mL}\cdot\text{kg}^{-1}\cdot\text{min}^{-1}$. This level was chosen because each one unit increase in $\dot{V}O_{2\max}$ is associated with a 16% predicted reduction in cancer mortality (10).

Association to cancer mortality risk. Our study population was at high risk for breast cancer. Therefore, we sought to further categorize change in $\dot{V}O_{2\max}$ by predicted risk reduction in cancer mortality. The relation between changes in fitness capacity to risk of cancer mortality was assessed by distributing change in $\dot{V}O_{2\max}$ across $1 \text{ mL}\cdot\text{kg}^{-1}\cdot\text{min}^{-1}$ categories via histogram plots at CONT, LOW, and HIGH groups, respectively. Each $1 \text{ mL}\cdot\text{kg}^{-1}\cdot\text{min}^{-1}$ category is associated with a 16% reduction in cancer mortality (10).

Statistical analysis. Sociodemographic factors (age, race, ethnicity, and education level) and baseline clinical characteristics (BMI and $\dot{V}O_{2\max}$) were compared using Fisher's exact test for categorical variables and one-way ANOVA for continuous variables. In addition, the number of exercise "nonresponders" was compared between groups by Fisher's exact test. Normality was tested by Shapiro–Wilk test to assess changes in $\dot{V}O_{2\max}$ from baseline to follow-up by groups. To estimate the probability density function of the associated risk reduction in cancer mortality, the normal density estimation curve was applied when the data are normally distributed (control group), whereas the kernel density estimation curve was used when the data were nonnormally distributed (intervention groups). Statistical analyses were conducted using SAS 9.4 (SAS Inc., Cary, NC). Statistical significance was set at an alpha level of 0.05.

RESULTS

Characteristics of the 125 participants are presented in Table 1. The mean age of participants was 34.6 ± 6.8 yr, and the mean BMI was 26.1 ± 6.0 $\text{kg}\cdot\text{m}^{-2}$. Participants were predominantly White (84.8%) and non-Hispanic (93.6%). We observed no significant between-group differences for age ($P = 0.68$), BMI ($P = 0.51$), $\dot{V}O_{2\max}$ preintervention ($P = 0.26$), race ($P = 0.67$), ethnicity ($P = 0.54$), and education level ($P = 0.37$). The WISER Sister study recruited women at high risk for breast cancer, and documentation of risk status by genetic testing or prediction model is given in Table 1. The comparison between individuals who withdrew from the study and individuals who completed the study has been previously discussed (35). Further, demographic and clinical characteristics were not different between responders and nonresponders in the intervention groups (data not shown).

Observed individual $\dot{V}O_{2\max}$ responses in the CONT (Fig. 1A), LOW (Fig. 1B), and HIGH (Fig. 1C) groups are presented. Although $2 \times \text{TE}$ (dashed line) is a relatively robust threshold

for the classification of "responders," it was calculated from tests approximately 6 months apart and likely represents change in the control group, rather than technical error or the coefficient of variation. Therefore, we chose CE (dotted line) as a better representative threshold. Participants were dichotomously classified as "responders" (Black) and "nonresponders" (White) based on the CE threshold of $1 \text{ mL}\cdot\text{kg}^{-1}\cdot\text{min}^{-1}$. There was a significant difference between groups in the number of nonresponders as exercise dose increased ($P < 0.01$): CONT ($n = 35$, 74.5%), LOW ($n = 11$, 28.2%), and HIGH ($n = 2$, 5.1%).

We next examined the adherence level of exercise "nonresponders." Overall, the average adherence level (adherence defined as proportion of total prescribed minutes completed within target heart rate range for the full study [22]) was 85% in the low-dose group and 81% in the high-dose group. We present individuals sorted by adherence level (Fig. 2). Given the adherence rate to the intervention, there were only 3 of 11 exercise "nonresponders" in the low-dose group that were nonadherent (<75% adherence), and their lack of adherence explains their lack of change in $\dot{V}O_{2\max}$ (Fig. 2A). However, 73% of "nonresponders" in the low-dose exercise group were true "nonresponders" as they did not improve their $\dot{V}O_{2\max}$ despite high adherence to the exercise prescription. Similarly, in the high-dose group, both "nonresponders" had high adherence (>90%) (Fig. 2B). The low-dose group maintained a significantly higher percentage of clinical nonresponders among participants who were adherent to the intervention (23.5%) compared with the high-dose group (5.6%) ($P = 0.04$). In addition, compared with age-appropriate general population $\dot{V}O_{2\max}$ norms, we observed that 57% of our study population was at good, excellent, or superior levels of fitness at baseline, and this increased to 66% by the follow-up exercise test (36).

Predicted cancer mortality risk reduction by dose of exercise was assessed (Fig. 3). In the CONT group (Fig. 3A), we observed a shift toward an increase in cancer mortality risk.

TABLE 1. Characteristics of 125 participants at study baseline.

Characteristic	Total Sample ($N = 125$)	Control ($n = 47$)	Low Dose ($n = 39$)	High Dose ($n = 39$)	<i>P</i>
Age (mean \pm SD), yr	34.6 ± 6.8	34.3 ± 7.6	35.4 ± 6.1	34.2 ± 6.4	0.68
BMI (mean \pm SD), $\text{kg}\cdot\text{m}^{-2}$	26.1 ± 6.0	26.7 ± 6.3	26.1 ± 6.2	25.2 ± 5.6	0.51
$\dot{V}O_{2\max}$ (mean \pm SD), $\text{mL}\cdot\text{kg}^{-1}\cdot\text{min}^{-1}$	32.6 ± 6.9	31.6 ± 7.4	32.2 ± 6.1	34.0 ± 7.1	0.26
Race, <i>n</i> (%)					0.67
White	106 (84.8)	39 (83.0)	32 (82.0)	35 (89.7)	
Non-White	19 (15.2)	8 (17.0)	7 (18.0)	4 (10.3)	
Ethnicity, <i>n</i> (%)					0.54
Hispanic	8 (6.4)	4 (8.5)	3 (7.7)	1 (2.6)	
Non-Hispanic	117 (93.6)	43 (91.5)	36 (92.3)	38 (97.4)	
Education level, <i>n</i> (%)					0.37
\leq High school degree	3 (2.4)	3 (6.4)	0	0	
Some college	29 (23.2)	12 (25.5)	9 (23.1)	8 (20.5)	
\geq College degree	93 (74.4)	32 (68.1)	30 (76.9)	31 (79.5)	
BRCA gene mutation status, <i>n</i> (%)					0.35
Positive	43 (34.4)	14 (29.8)	16 (41.0)	13 (33.3)	
Negative	11 (8.8)	7 (14.9)	1 (2.6)	3 (7.7)	
Not tested	71 (56.8)	26 (55.3)	22 (56.4)	23 (59.0)	
>18% lifetime risk of developing breast cancer, <i>n</i> (%)					
Gail ^a	50 (40.0)	18 (69.2)	18 (85.7)	14 (82.4)	0.40
Claus ^b	83 (66.4)	30 (65.2)	27 (69.2)	26 (68.4)	0.91

^aA Gail score is not calculated for women below the age of 35 yr.

^bA Claus score is not calculated for women who lack female first and/or degree relatives with breast cancer.

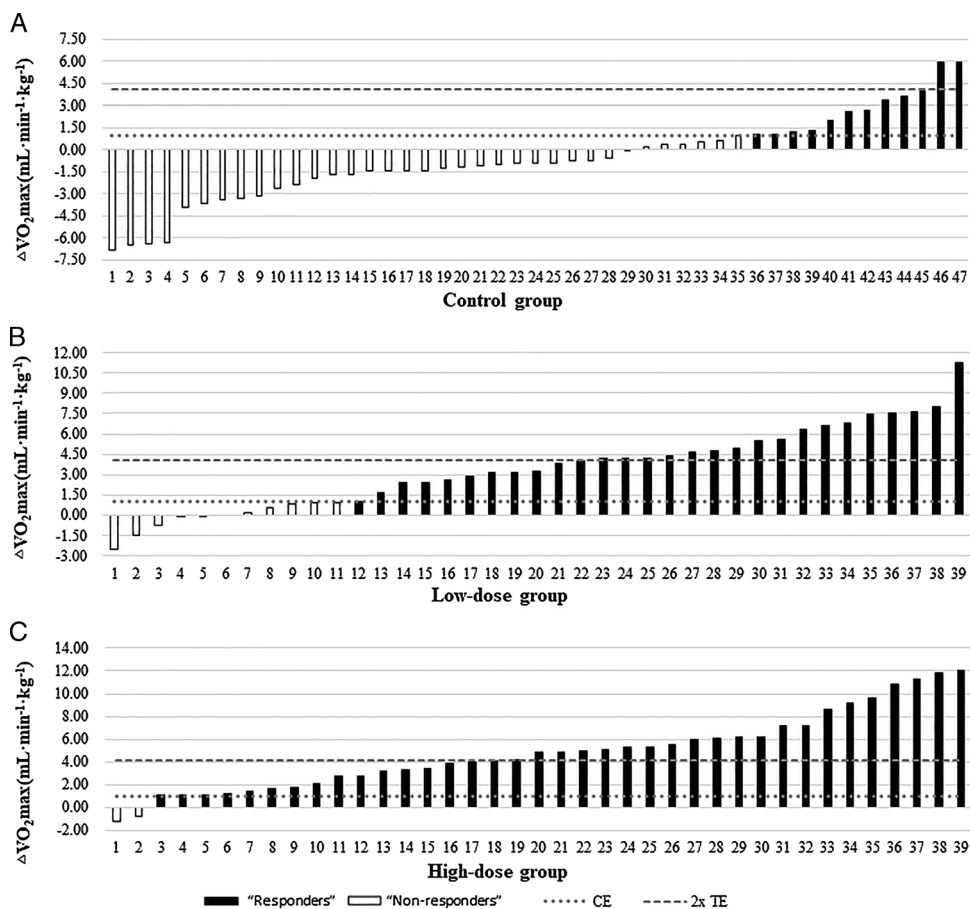


FIGURE 1—A, Observed individual $\dot{V}O_{2\max}$ response in CONT group and CE threshold-based dichotomous classification of responders and nonresponders. B, Observed individual $\dot{V}O_{2\max}$ response in low-dose group and CE threshold-based dichotomous classification of responders and nonresponders. C, Observed individual $\dot{V}O_{2\max}$ response in high-dose group and CE threshold-based dichotomous classification of responders and nonresponders.

In the low-dose (Fig. 3B) and high-dose groups (Fig. 3C), the distribution of cancer mortality risk shifted toward a decrease in cancer risk (peak kernel density estimation fitted between -80% and -64%). A greater number of participants reduced their cancer mortality risk in the high-dose group (94.9%) compared with the low-dose group (87.2%), although the difference was not significant ($P = 0.43$).

DISCUSSION

The current study investigated change in $\dot{V}O_{2\max}$ at the individual level after a 6-month aerobic exercise intervention among premenopausal women at elevated breast cancer risk. We found that exercise “nonresponders” occurred in both low-dose and high-dose groups independent of relative adherence to the exercise intervention. We observed that increasing the volume of exercise dose (via exercise duration) significantly decreased the number of “nonresponders.” Nonresponders were classified by clinical effectiveness, which was defined as the change in $\dot{V}O_{2\max}$ associated with reduced risk from cancer mortality. Increasing the volume of exercise dose did not provide additional benefit for predicted decreases in cancer mortality risk.

In addition to the aforementioned findings, we addressed an important methodological issue in terms of the classification of “responders” and “nonresponders.” We adopted the CE threshold for the classification of individual responses. We observed more “responders” to exercise with CE threshold compared with $2 \times TE$. This is because $2 \times TE$ is a statistical measure referring to a change that is equal to two SE of the mean from the first test, but our exercise tests were 6 months apart. Thus, $2 \times TE$ may overestimate technical error in measurement, and instead reflect actual physiological change over 6 months. As a population at high risk of breast cancer, our participants face a serious threat of death from cancer. Thus, our selection of CE, which relates changes in $\dot{V}O_{2\max}$ to cancer mortality risk, is reasonable and clearly fits clinical meaningfulness.

The WISER Sister study has measured several biomarkers of breast cancer risk. Variables related to body composition, estrogen sensitive breast tissue, estrogen and progesterone levels, adipokine levels, and cytokine levels have all been reported previously (23–25,35). Exercise training-induced changes in these variables have not been linked to clinical outcomes. Therefore, change in $\dot{V}O_{2\max}$ is presently the strongest biomarker available from our data set for predicting how change after an exercise intervention can affect future cancer risk. Although we chose to classify “responders” and “nonresponders”

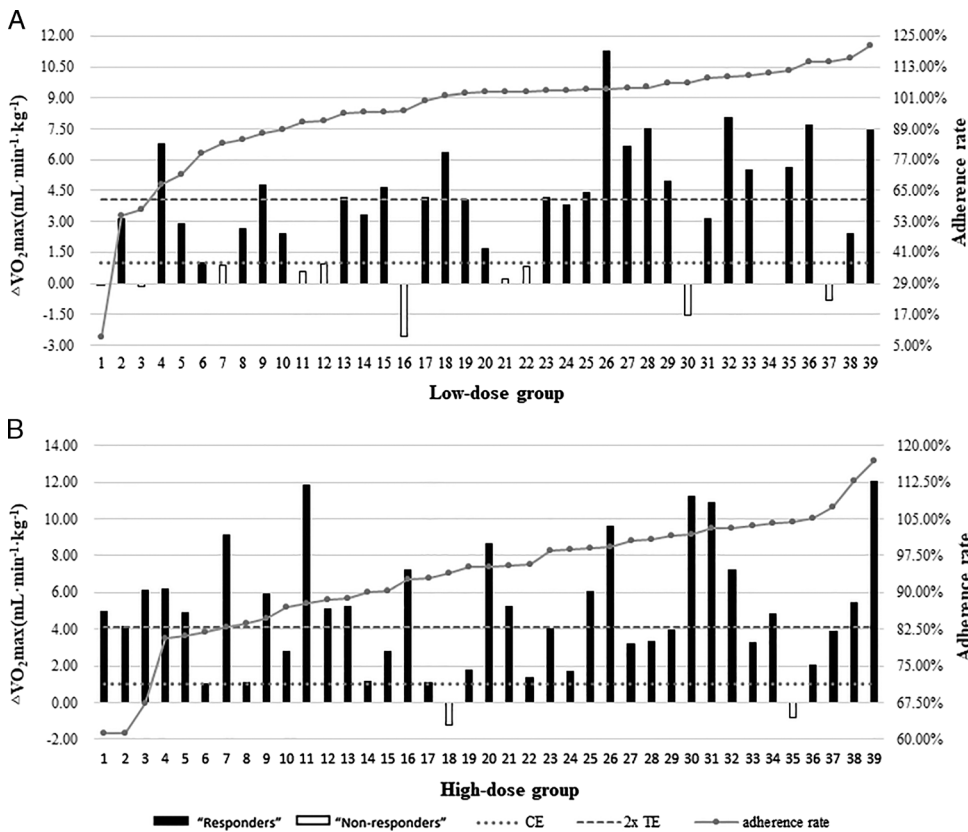


FIGURE 2—A, Observed individual $\dot{V}O_{2max}$ responses in low-dose group sorted by adherence rate and CE threshold-based dichotomous classification of responders and nonresponders. B, Observed individual $\dot{V}O_{2max}$ responses in high-dose group sorted by adherence rate and CE threshold-based dichotomous classification of responders and nonresponders.

to exercise based on $\dot{V}O_{2max}$, it is still likely that individuals who did not improve their $\dot{V}O_{2max}$ may have experienced other benefits from exercise.

Findings of the current study provide useful implications for exercise prescription decision making. As our study population

is at high risk of breast cancer, determining an appropriate amount of exercise to reduce the associated cancer mortality is important. Although we observed that a higher exercise dose is more beneficial in improving $\dot{V}O_{2max}$, it did not significantly increase the number of participants with a predicted

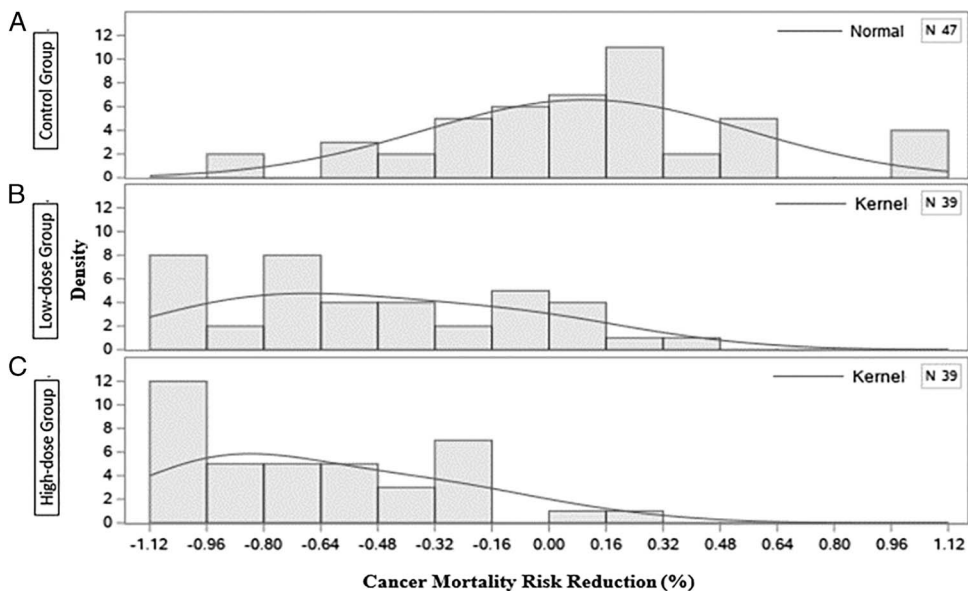


FIGURE 3—A, Change in $\dot{V}O_{2max}$ -associated cancer mortality risk reduction in CONT group. B, Change in $\dot{V}O_{2max}$ -associated cancer mortality risk reduction in low-dose group. C, Change in $\dot{V}O_{2max}$ -associated cancer mortality risk reduction in high-dose group.

reduction in cancer mortality risk. This may indicate that $150 \text{ min}\cdot\text{wk}^{-1}$ is sufficient to control the risk of cancer mortality for the majority of people. Yet, for those that fail to increase $\dot{V}O_{2\text{max}}$ at $150 \text{ min}\cdot\text{wk}^{-1}$, these individuals may benefit from increasing their exercise duration. Overall, our observations are consistent with current U.S. Department of Health and Human Services and American College of Sports Medicine guidance: for substantial health benefits, adults should do at least 150 to 300 $\text{min}\cdot\text{wk}^{-1}$ of moderate-intensity aerobic physical activity (37).

Identification of exercise “nonresponders” can provide useful information for developing personalized medicine. As $\dot{V}O_{2\text{max}}$ is one of the best clinical predictors of all-cause mortality (second only to smoking status), doing $\dot{V}O_{2\text{max}}$ testing after an exercise prescription would allow a clinician to determine efficacy of the exercise prescription (38). Similar to thyroid dysfunction, a physician prescribes a drug and retests thyroid levels for appropriate response. Just as a physician would modulate the drug prescription based on follow-up testing, so too can an exercise physiologist. Our study has provided the exercise physiologist with evidence to further prescribe additional duration of aerobic exercise at the same intensity. Future studies should determine whether increasing exercise intensity will decrease rate of nonresponders. It appears that high-intensity interventions also result in nonresponders (28,31); thus, more work on exercise prescriptions is necessary to understand how to dose and prescribe exercise modalities.

Nonpharmacological cancer prevention methods are particularly important for our study population as, currently, effective preventative measures are limited to prophylactic bilateral mastectomies, bilateral oophorectomies, and selective estrogen receptor modulators (39,40). These options are expensive and carry their own risks in the short- and long-term settings. Although not necessarily a replacement for these therapies, aerobic exercise can reduce breast cancer risk, delay onset, and prove to be a valuable preventative measure in high-risk individuals (41–45).

There are two major strengths of our study: 1) high exercise adherence and 2) consistency in exercise intensity and exercise type between groups. Adherence to the exercise intervention is an important factor to assess the accuracy of classifying exercise nonresponders. Without accounting for individual exercise adherence, the density of clinical nonresponders can be overestimated. Our study rigorously assessed adherence to the prescribed exercise intervention using both self-reported logs and objective monitoring. We observed some women did not improve their fitness capacity despite high adherence levels in both exercise intervention groups, which were true clinical “nonresponders” to exercise. Also, this secondary analysis examines modulation of the frequency, intensity, time, and type principle of exercise physiology. To assess the volume of exercise via an increase in time (150 or $300 \text{ min}\cdot\text{wk}^{-1}$), we strictly controlled the exercise intensity and type between groups.

Although the current study explored CE and $2\times$ TE for the classification of “responders” and “nonresponders,” it is important to note that there are several other different thresholds

that can be used (46). For instance, a fixed proportion of participants with the lowest training response were classified as “nonresponders” in some studies, regardless of the magnitude of change in the outcome of interest (47,48). Other groups have set a threshold of 0.5 metabolic equivalents of task (0.5 MET, equals to $1.75 \text{ mL}\cdot\text{kg}^{-1}\cdot\text{min}^{-1} \dot{V}O_{2\text{max}}$) and found three exercise “nonresponders” among 29 healthy, recreationally active participants (10.3%) after 4 wk of structured exercise training (33). Furthermore, some pointed out a concept of the smallest worthwhile change as a meaningful value to define their threshold in identifying “nonresponders,” which is quite similar to the definition of CE we used (32,49). It is worth mentioning that Swinton et al. (49) recommended choosing the smallest worthwhile change that is lower than the expected change for most individuals. This argument is consistent with our study methodological issue that the value of CE we adopted is lower than the $2\times$ TE we calculated. Using a different threshold will significantly affect the picture of individuals classified as responders and nonresponders (46). As there remains lack of consistency as to how “nonresponders” should be defined, the threshold should be carefully considered in the research context, e.g., what the exercise is prescribed for, who is the main population, how long is the intervention period, and what are desired outcomes.

Other limitations include that our study population was predominantly educated non-Hispanic White women. The incidence of breast cancer is similar between non-Hispanic White and Black women, yet the mortality rate among Black women due to breast cancer is 42% higher (50). Although our findings provide promise, the implications of our findings exclude populations from diverse racial and ethnic backgrounds. Future studies should assess the association between the change in cancer-specific mortality in a larger cohort of more diverse women with greater disease-specific mortality end points.

In conclusion, this study supports the application of a CE threshold to categorize clinical exercise “nonresponders” and demonstrates a dose–response effect via time spent exercising on decreasing the number of “nonresponders” to exercise. The $150\text{-min}\cdot\text{wk}^{-1}$ exercise volume may also be sufficient to decrease the predicted risk of cancer mortality for women at high risk of developing breast cancer. Based on our findings, future studies should assess the modulation of dosed exercise prescriptions (frequency, intensity, time, and type).

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REFERENCES

- Coughlin SS, Ekwueme DU. Breast cancer as a global health concern. *Cancer Epidemiol*. 2009;33(5):315–8.
- Ghoncheh M, Pournamdar Z, Salehiniya H. Incidence and mortality and epidemiology of breast cancer in the world. *Asian Pac J Cancer Prev*. 2016;17(S3):43–6.
- Ghoncheh M, Momenimovahed Z, Salehiniya H. Epidemiology, incidence and mortality of breast cancer in Asia. *Asian Pac J Cancer Prev*. 2016;17(S3):47–52.
- Bray F, Ferlay J, Soerjomataram I, Siegel RL, Torre LA, Jemal A. Global cancer statistics 2018: GLOBOCAN estimates of incidence and mortality worldwide for 36 cancers in 185 countries. *CA Cancer J Clin*. 2018;68(6):394–424.
- Shulman LN, Willett W, Sievers A, Knaul FM. Breast cancer in developing countries: opportunities for improved survival. *J Oncol*. 2010;2010:1–6.
- DeSantis CE, Ma J, Goding Sauer A, Newman LA, Jemal A. Breast cancer statistics, 2017, racial disparity in mortality by state. *CA Cancer J Clin*. 2017;67(6):439–48.
- Courneya KS, Mackey JR, Bell GJ, Jones LW, Field CJ, Fairey AS. Randomized controlled trial of exercise training in postmenopausal breast cancer survivors: cardiopulmonary and quality of life outcomes. *J Clin Oncol*. 2003;21(9):1660–8.
- McNeely ML, Campbell KL, Rowe BH, Klassen TP, Mackey JR, Courneya KS. Effects of exercise on breast cancer patients and survivors: a systematic review and meta-analysis. *CMAJ*. 2006;175(1):34–41.
- Furmaniak AC, Menig M, Markes MH. Exercise for women receiving adjuvant therapy for breast cancer. *Cochrane Database Syst Rev*. 2016;9:CD005001.
- Imboden MT, Harber MP, Whaley MH, et al. The association between the change in directly measured cardiorespiratory fitness across time and mortality risk. *Prog Cardiovasc Dis*. 2019;62(2):157–62.
- Holick CN, Newcomb PA, Trentham-Dietz A, et al. Physical activity and survival after diagnosis of invasive breast cancer. *Cancer Epidemiol Biomarkers Prev*. 2008;17(2):379–86.
- Irwin ML, McTiernan A, Manson JE, et al. Physical activity and survival in postmenopausal women with breast cancer: results from the women’s health initiative. *Cancer Prev Res*. 2011;4(4):522–9.
- Beasley JM, Kwan ML, Chen WY, et al. Meeting the physical activity guidelines and survival after breast cancer: findings from the after breast cancer pooling project. *Breast Cancer Res Treat*. 2012;131(2):637–43.
- Bouchard C, Rankinen T. Individual differences in response to regular physical activity. *Med Sci Sports Exerc*. 2001;33(6 Suppl):S446–51; discussion S452–3.
- Hopkins WG. Precision of the estimate of a subject’s true value (Excel spreadsheet). In: *A New View of Statistics sports sci org: Internet Society for Sport Science*; 2000. Available from: <http://www.sportsci.org/resource/stats/>. Accessed February 4, 2020.
- Buford TW, Roberts MD, Church TS. Toward exercise as personalized medicine. *Sports Med*. 2013;43(3):157–65.
- Gurd BJ, Giles MD, Bonafiglia JT, et al. Incidence of nonresponse and individual patterns of response following sprint interval training. *Appl Physiol Nutr Metab*. 2016;41(3):229–34.
- Williamson PJ, Atkinson G, Batterham AM. Inter-individual responses of maximal oxygen uptake to exercise training: a critical review. *Sports Med*. 2017;47(8):1501–13.
- Scharhag-Rosenberger F, Walitzek S, Kindermann W, Meyer T. Differences in adaptations to 1 year of aerobic endurance training: individual patterns of nonresponse. *Scand J Med Sci Sports*. 2012;22(1):113–8.
- Pickering C, Kiely J. Do non-responders to exercise exist and if so, what should we do about them? *Sports Med*. 2019;49(1):1–7.
- Senn S. Individual response to treatment: is it a valid assumption? *BMJ*. 2004;329(7472):966–8.
- Schmitz KH, Williams NI, Kontos D, et al. Women in Steady Exercise Research (WISER) Sister: study design and methods. *Contemp Clin Trials*. 2015;41:17–30.
- Schmitz KH, Williams NI, Kontos D, et al. Dose–response effects of aerobic exercise on estrogen among women at high risk for breast cancer: a randomized controlled trial. *Breast Cancer Res Treat*. 2015;154(2):309–18.
- Sturgeon K, Digiovanni L, Good J, et al. Exercise-induced dose–response alterations in adiponectin and leptin levels are dependent on body fat changes in women at risk for breast cancer. *Cancer Epidemiol Biomarkers Prev*. 2016;25(8):1195–200.
- Brown JC, Kontos D, Schnall MD, Wu S, Schmitz KH. The dose–response effects of aerobic exercise on body composition and breast tissue among women at high risk for breast cancer: a randomized trial. *Cancer Prev Res (Phila)*. 2016;9(7):581–8.
- Bruce RA, Kusumi F, Hosmer D. Maximal oxygen intake and nomographic assessment of functional aerobic impairment in cardiovascular disease. *Am Heart J*. 1973;85(4):546–62.
- Bouchard C, Blair SN, Church TS, et al. Adverse metabolic response to regular exercise: is it a rare or common occurrence? *PLoS One*. 2012;7(5):e37887.
- Bonafiglia JT, Rotundo MP, Whittall JP, Scribbans TD, Graham RB, Gurd BJ. Inter-individual variability in the adaptive responses to endurance and sprint interval training: a randomized crossover study. *PLoS One*. 2016;11(12):e0167790.
- Álvarez C, Ramírez-Campillo R, Ramírez-Vélez R, Izquierdo M. Prevalence of non-responders for glucose control markers after 10 weeks of high-intensity interval training in adult women with higher and lower insulin resistance. *Front Physiol*. 2017;8:479.
- de Lannoy L, Clarke J, Stotz PJ, Ross R. Effects of intensity and amount of exercise on measures of insulin and glucose: analysis of inter-individual variability. *PLoS One*. 2017;12(5):e0177095.
- Astorino TA, deRevere J, Anderson T, et al. Change in $\dot{V}O_{2max}$ and time trial performance in response to high-intensity interval training prescribed using ventilatory threshold. *Eur J Appl Physiol*. 2018;118(9):1811–20.
- Hopkins WG. Measures of reliability in sports medicine and science. *Sports Med*. 2000;30(1):1–15.
- Bonafiglia JT, Nelms MW, Preobrazenski N, et al. Moving beyond threshold-based dichotomous classification to improve the accuracy in classifying non-responders. *Physiol Rep*. 2018;6(22):e13928.
- Ashcroft R. What is clinical effectiveness? *Stud Hist Philos Biol Biomed Sci*. 2002;33(2):219–33.
- Haley JS, Hibler EA, Zhou S, Schmitz KH, Sturgeon KM. Dose-dependent effect of aerobic exercise on inflammatory biomarkers in a randomized controlled trial of women at high risk of breast cancer. *Cancer*. 2020;126(2):329–36.
- Cooper I. *Physical Fitness Assessments and Norms for Adults and Law Enforcement*. Dallas (TX): The Cooper Institute; 2007.
- Garber CE, Blissmer B, Deschenes MR, et al, American College of Sports Medicine. American College of Sports Medicine Position Stand: quantity and quality of exercise for developing and maintaining cardiorespiratory, musculoskeletal, and neuromotor fitness in apparently healthy adults: guidance for prescribing exercise. *Med Sci Sports Exerc*. 2011;43(7):1334–59.
- Ladenvall P, Persson CU, Mandalenakis Z, et al. Low aerobic capacity in middle-age men associated with increased mortality rates during 45 years of follow-up. *Eur J Prev Cardiol*. 2016;23(14):1557–64.
- Nelson HD, Pappas M, Zakher B, Mitchell JP, Okinaka-Hu L, Fu R. Risk assessment, genetic counseling, and genetic testing for BRCA-related cancer in women: a systematic review to update the U.S. Preventive Services Task Force recommendation. *Ann Intern Med*. 2014;160(4):255–66.
- Grann VR, Patel PR, Jacobson JS, et al. Comparative effectiveness of screening and prevention strategies among BRCA1/2-affected mutation carriers. *Breast Cancer Res Treat*. 2011;125(3):837–47.
- Thune I, Brenn T, Lund E, Gaard M. Physical activity and the risk of breast cancer. *N Engl J Med*. 1997;336(18):1269–75.

42. Rockhill B, Willett WC, Hunter DJ, Manson JE, Hankinson SE, Colditz GA. A prospective study of recreational physical activity and breast cancer risk. *Arch Intern Med.* 1999;159(19):2290–6.
43. Pijpe A, Manders P, Brohet RM, et al. Physical activity and the risk of breast cancer in BRCA1/2 mutation carriers. *Breast Cancer Res Treat.* 2010;120(1):235–44.
44. Wu Y, Zhang D, Kang S. Physical activity and risk of breast cancer: a meta-analysis of prospective studies. *Breast Cancer Res Treat.* 2013;137(3):869–82.
45. Lammert J, Lubinski J, Gronwald J, et al. Physical activity during adolescence and young adulthood and the risk of breast cancer in BRCA1 and BRCA2 mutation carriers. *Breast Cancer Res Treat.* 2018;169(3):561–71.
46. Swank AM, Horton J, Fleg JL, et al. Modest increase in peak VO₂ is related to better clinical outcomes in chronic heart failure patients: results from heart failure and a controlled trial to investigate outcomes of exercise training. *Circ Heart Fail.* 2012;5(5):579–85.
47. Vollaard NB, Constantin-Teodosiu D, Fredriksson K, et al. Systematic analysis of adaptations in aerobic capacity and submaximal energy metabolism provides a unique insight into determinants of human aerobic performance. *J Appl Physiol (1985).* 2009;106(5):1479–86.
48. Timmons JA, Knudsen S, Rankinen T, et al. Using molecular classification to predict gains in maximal aerobic capacity following endurance exercise training in humans. *J Appl Physiol.* 2010;108(6):1487–96.
49. Swinton PA, Hemingway BS, Saunders B, Gualano B, Dolan E. A statistical framework to interpret individual response to intervention: paving the way for personalized nutrition and exercise prescription. *Front Nutr.* 2018;5:41.
50. DeSantis CE, Fedewa SA, Goding Sauer A, Kramer JL, Smith RA, Jemal A. Breast cancer statistics, 2015: convergence of incidence rates between Black and White women. *CA Cancer J Clin.* 2016;66(1):31–42.